# Natural Language Models and Interfaces

BSc Artificial Intelligence

Lecturer: Wilker Aziz Institute for Logic, Language, and Computation

2020, week 5

# Data sparsity

We have so far dealt with categorical models using tabular CPDs

- we've encountered problems for maximum likelihood estimation due to data sparsity
- ▶ large n-grams lead to large tables  $O(v^n)$
- even the emission distributions of HMMs had to be smoothed
- smoothing techniques can be rather brittle

Are there principled ways to tackle data sparsity?

# Data sparsity

We have so far dealt with categorical models using tabular CPDs

- we've encountered problems for maximum likelihood estimation due to data sparsity
- ▶ large n-grams lead to large tables  $O(v^n)$
- even the emission distributions of HMMs had to be smoothed
- smoothing techniques can be rather brittle

Are there principled ways to tackle data sparsity?

Let's check a running example based on *sentiment classification* 

#### Example: sentiment classification

How can we identify whether a sentence is positive or negative towards a subject?

- ► This movie is slow and repetitive, clearly the direction was careless and the production cheap.
- ► The movie is quite funny, its spiced humor makes it very interesting.

#### Example: sentiment classification

How can we identify whether a sentence is positive or negative towards a subject?

- ► This movie is slow and repetitive, clearly the direction was careless and the production cheap.
- ► The movie is quite funny, its spiced humor makes it very interesting.

If y is the sentiment of some text x, then we would like to compute  $P_{Y\mid X}(y\mid x)$ , but

- x is very sparse!
- how could we possibly parameterise it?

One way to model  $P_{Y|X}(y|x)$  is to model it directly, i.e. directly parameterising a distribution over  $\mathcal{X} \times \mathcal{Y}$ .

One way to model  $P_{Y|X}(y|x)$  is to model it directly, i.e. directly parameterising a distribution over  $\mathcal{X} \times \mathcal{Y}$ .

Another way is to infer this conditional from a joint distribution  $P_{XY}(x,y)$ .

One way to model  $P_{Y|X}(y|x)$  is to model it directly, i.e. directly parameterising a distribution over  $\mathcal{X} \times \mathcal{Y}$ .

Another way is to infer this conditional from a joint distribution  $P_{XY}(x,y)$ . But how is this useful?

One way to model  $P_{Y|X}(y|x)$  is to model it directly, i.e. directly parameterising a distribution over  $\mathcal{X} \times \mathcal{Y}$ .

Another way is to infer this conditional from a joint distribution  $P_{XY}(x,y)$ . But how is this useful?

Because we have an opportunity to make conditional independence assumptions and directly parameterise distributions over smaller sample spaces instead!

#### Feature functions

Suppose we identify a number of words which typically express sentiment, let's call them *features* 

Class	Features (or attributes)
Negative	cheap, slow, repetitive, careless, awful, bad,
Positive	funny, spiced, interesting, awesome, good,

and let's say we have a vocabulary  ${\mathcal F}$  of  ${\it v}$  such sentiment words

#### Feature functions

Suppose we identify a number of words which typically express sentiment, let's call them *features* 

Class	Features (or attributes)
Negative	cheap, slow, repetitive, careless, awful, bad,
Positive	funny, spiced, interesting, awesome, good,

and let's say we have a vocabulary  ${\mathcal F}$  of v such sentiment words

#### Then for an example

x= 'This film is fun though shy on the action' with sentiment y=+, let us retain only the sentiment words:

 $\langle \text{`fun', 'shy', 'though', 'action'} \rangle = \\ \text{sentwords('This film is fun though shy on the action')}$ 

#### Feature functions

Suppose we identify a number of words which typically express sentiment, let's call them *features* 

Class	Features (or attributes)
Negative	cheap, slow, repetitive, careless, awful, bad,
Positive	funny, spiced, interesting, awesome, good,

and let's say we have a vocabulary  ${\mathcal F}$  of v such sentiment words

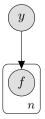
#### Then for an example

x= 'This film is fun though shy on the action' with sentiment y=+, let us retain only the sentiment words:

- ⟨'fun', 'shy', 'though', 'action'⟩ = sentwords('This film is fun though shy on the action')
- we will denote this by a random pair  $(Y, \langle F_1, \dots, F_n \rangle)$

# Naive Bayes Classifiers

In a naive Bayes classifier, we assume the *class* y generates the features  $\langle f_1, \dots, f_n \rangle$ 

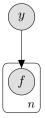


That is, we assume features to be conditionally independent given a class

$$P_{YF_1^n}(y,\langle f_1,\ldots,f_n\rangle)=$$

## Naive Bayes Classifiers

In a naive Bayes classifier, we assume the *class* y generates the features  $\langle f_1, \ldots, f_n \rangle$ 

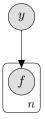


That is, we assume features to be conditionally independent given a class

$$P_{YF_1^n}(y,\langle f_1,\ldots,f_n\rangle)=P_Y(y)\times$$

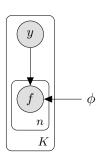
## Naive Bayes Classifiers

In a naive Bayes classifier, we assume the *class* y generates the features  $\langle f_1, \ldots, f_n \rangle$ 

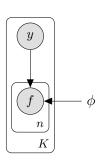


That is, we assume features to be conditionally independent given a class

$$P_{YF_1^n}(y,\langle f_1,\ldots,f_n\rangle) = P_Y(y) \times \prod_{i=1}^n P_{F|Y}(f_i|y)$$



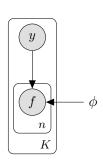
Suppose a dataset of labelled examples  $\mathcal{D} = \left\{y^{(k)}, \langle f_1^{(k)}, \dots, f_{n_k}^{(k)} \rangle \right\}_{k=1}^K \text{ and a vocabulary of } v \text{ features}$ 



Suppose a dataset of labelled examples  $\mathcal{D} = \left\{y^{(k)}, \langle f_1^{(k)}, \dots, f_{n_k}^{(k)} \rangle \right\}_{k=1}^K \text{ and a vocabulary of } v \text{ features}$ 

The generative story for each training instance:

$$Y \sim \mathcal{U}(1/2)$$
 for  $i = 1, \dots, n$  
$$F_i | y \sim \operatorname{Cat}(\phi_1^{(y)}, \dots, \phi_v^{(y)})$$



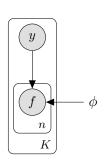
Suppose a dataset of labelled examples  $\mathcal{D} = \left\{y^{(k)}, \langle f_1^{(k)}, \dots, f_{n_k}^{(k)} \rangle \right\}_{k=1}^K \text{ and a vocabulary of } v \text{ features}$ 

The generative story for each training instance:

$$Y \sim \mathcal{U}(1/2)$$
 for  $i = 1, \dots, n$  
$$F_i | y \sim \operatorname{Cat}(\phi_1^{(y)}, \dots, \phi_v^{(y)})$$

Thus, for one training instance

$$P_{YF_1^n}(y,\langle f_1,\ldots,f_n\rangle) = P_Y(y) \times \prod_{i=1}^n P_{F|Y}(f_i|y)$$



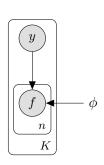
Suppose a dataset of labelled examples  $\mathcal{D} = \left\{y^{(k)}, \langle f_1^{(k)}, \dots, f_{n_k}^{(k)} \rangle \right\}_{k=1}^K \text{ and a vocabulary of } v \text{ features}$ 

The generative story for each training instance:

$$Y \sim \mathcal{U}(1/2)$$
 for  $i = 1, \dots, n$  
$$F_i | y \sim \operatorname{Cat}(\phi_1^{(y)}, \dots, \phi_v^{(y)})$$

Thus, for one training instance

$$P_{YF_1^n}(y, \langle f_1, \dots, f_n \rangle) = P_Y(y) \times \prod_{i=1}^n P_{F|Y}(f_i|y)$$
$$= \mathcal{U}(1/2) \prod_{i=1}^n \operatorname{Cat}(f_i|\phi^{(y)})$$



Suppose a dataset of labelled examples  $\mathcal{D} = \left\{y^{(k)}, \langle f_1^{(k)}, \dots, f_{n_k}^{(k)} \rangle \right\}_{k=1}^K \text{ and a vocabulary of } v \text{ features}$ 

The generative story for each training instance:

$$Y \sim \mathcal{U}(1/2)$$
 for  $i = 1, \dots, n$  
$$F_i | y \sim \operatorname{Cat}(\phi_1^{(y)}, \dots, \phi_v^{(y)})$$

6

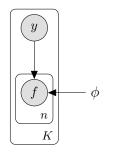
Thus, for one training instance

$$P_{YF_1^n}(y,\langle f_1,\ldots,f_n\rangle) = P_Y(y) \times \prod_{i=1}^n P_{F|Y}(f_i|y)$$
$$= \mathcal{U}(1/2) \prod_{i=1}^n \operatorname{Cat}(f_i|\phi^{(y)}) \propto \prod_{i=1}^n \phi_{f_i}^{(y)}$$

The log-likelihood of the data is proportional to

$$\log P(\mathcal{D}|\phi) = \prod_{k=1}^{K} P_{Y|F_1^n}(y_k, \langle f_1^{(k)}, \dots, f_{n_k}^{(k)} \rangle | \phi)$$

$$\propto \sum_{k=1}^{K} \sum_{i=1}^{n} \log P_{F|Y}(f_i^{(k)} | y^{(k)}, \phi)$$

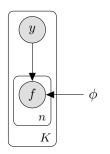


How many parameters?

The log-likelihood of the data is proportional to

$$\log P(\mathcal{D}|\phi) = \prod_{k=1}^{K} P_{Y|F_1^n}(y_k, \langle f_1^{(k)}, \dots, f_{n_k}^{(k)} \rangle | \phi)$$

$$\propto \sum_{k=1}^{K} \sum_{i=1}^{n} \log P_{F|Y}(f_i^{(k)} | y^{(k)}, \phi)$$

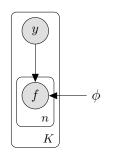


How many parameters?  $2 \times v$ 

The log-likelihood of the data is proportional to

$$\log P(\mathcal{D}|\phi) = \prod_{k=1}^{K} P_{Y|F_1^n}(y_k, \langle f_1^{(k)}, \dots, f_{n_k}^{(k)} \rangle | \phi)$$

$$\propto \sum_{k=1}^{K} \sum_{i=1}^{n} \log P_{F|Y}(f_i^{(k)} | y^{(k)}, \phi)$$



How many parameters?  $2 \times v$ 

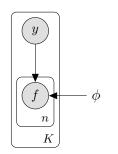
MLE

$$\phi_f^{(y)} =$$

The log-likelihood of the data is proportional to

$$\log P(\mathcal{D}|\phi) = \prod_{k=1}^{K} P_{Y|F_1^n}(y_k, \langle f_1^{(k)}, \dots, f_{n_k}^{(k)} \rangle | \phi)$$

$$\propto \sum_{k=1}^{K} \sum_{i=1}^{n} \log P_{F|Y}(f_i^{(k)} | y^{(k)}, \phi)$$



How many parameters?  $2 \times v$ 

MLE

$$\phi_f^{(y)} = \frac{\text{count}_{YF}(y, f)}{\text{count}_Y(y)}$$

For some new example with features  $\langle f_1, \ldots, f_n \rangle$ , we can predict its class easily by solving a maximisation problem

$$y^* = \underset{y}{\operatorname{argmax}} P_{Y|F_1^n}(y|\langle f_1, \dots, f_n \rangle)$$

For some new example with features  $\langle f_1, \ldots, f_n \rangle$ , we can predict its class easily by solving a maximisation problem

$$y^* = \underset{y}{\operatorname{argmax}} P_{Y|F_1^n}(y|\langle f_1, \dots, f_n \rangle)$$

$$= \underset{y}{\operatorname{argmax}} \frac{P_Y(y)P_{F_1^n|Y}(\langle f_1, \dots, f_n \rangle|y)}{P_{F_1^n}(\langle f_1, \dots, f_n \rangle)}$$

Bayes rule

For some new example with features  $\langle f_1, \ldots, f_n \rangle$ , we can predict its class easily by solving a maximisation problem

$$y^* = \underset{y}{\operatorname{argmax}} \ P_{Y|F_1^n}(y|\langle f_1, \dots, f_n \rangle)$$

$$= \underset{y}{\operatorname{argmax}} \ \frac{P_Y(y)P_{F_1^n|Y}(\langle f_1, \dots, f_n \rangle|y)}{P_{F_1^n}(\langle f_1, \dots, f_n \rangle)}$$

$$= \underset{y}{\operatorname{argmax}} \ P_Y(y)P_{F_1^n|Y}(\langle f_1, \dots, f_n \rangle|y)$$
Proportionality

For some new example with features  $\langle f_1, \ldots, f_n \rangle$ , we can predict its class easily by solving a maximisation problem

$$\begin{split} y^{\star} &= \underset{y}{\operatorname{argmax}} & P_{Y|F_{1}^{n}}(y|\langle f_{1}, \dots, f_{n} \rangle) \\ &= \underset{y}{\operatorname{argmax}} & \frac{P_{Y}(y)P_{F_{1}^{n}|Y}(\langle f_{1}, \dots, f_{n} \rangle|y)}{P_{F_{1}^{n}}(\langle f_{1}, \dots, f_{n} \rangle|y)} & \text{Bayes rule} \\ &= \underset{y}{\operatorname{argmax}} & P_{Y}(y)P_{F_{1}^{n}|Y}(\langle f_{1}, \dots, f_{n} \rangle|y) & \text{Proportionality} \\ &= \underset{y}{\operatorname{argmax}} & P_{F_{1}^{n}|Y}(\langle f_{1}, \dots, f_{n} \rangle|y) & \text{Proportionality} \end{split}$$

For some new example with features  $\langle f_1, \ldots, f_n \rangle$ , we can predict its class easily by solving a maximisation problem

$$\begin{split} y^{\star} &= \underset{y}{\operatorname{argmax}} \ P_{Y|F_{1}^{n}}(y|\langle f_{1}, \dots, f_{n} \rangle) \\ &= \underset{y}{\operatorname{argmax}} \ \frac{P_{Y}(y)P_{F_{1}^{n}|Y}(\langle f_{1}, \dots, f_{n} \rangle|y)}{P_{F_{1}^{n}}(\langle f_{1}, \dots, f_{n} \rangle)} \qquad \qquad \text{Bayes rule} \\ &= \underset{y}{\operatorname{argmax}} \ P_{Y}(y)P_{F_{1}^{n}|Y}(\langle f_{1}, \dots, f_{n} \rangle|y) \qquad \qquad \text{Proportionality} \\ &= \underset{y}{\operatorname{argmax}} \ P_{F_{1}^{n}|Y}(\langle f_{1}, \dots, f_{n} \rangle|y) \qquad \qquad \text{Proportionality} \\ &= \underset{y}{\operatorname{argmax}} \ \prod_{i=1}^{n} P_{F|Y}(f_{i}|y) \qquad \qquad \text{Conditional independence} \end{split}$$

For some new example with features  $\langle f_1, \ldots, f_n \rangle$ , we can predict its class easily by solving a maximisation problem

$$\begin{split} y^{\star} &= \underset{y}{\operatorname{argmax}} & P_{Y|F_{1}^{n}}(y|\langle f_{1}, \dots, f_{n} \rangle) \\ &= \underset{y}{\operatorname{argmax}} & \frac{P_{Y}(y)P_{F_{1}^{n}|Y}(\langle f_{1}, \dots, f_{n} \rangle|y)}{P_{F_{1}^{n}}(\langle f_{1}, \dots, f_{n} \rangle|y)} & \text{Bayes rule} \\ &= \underset{y}{\operatorname{argmax}} & P_{Y}(y)P_{F_{1}^{n}|Y}(\langle f_{1}, \dots, f_{n} \rangle|y) & \text{Proportionality} \\ &= \underset{y}{\operatorname{argmax}} & P_{F_{1}^{n}|Y}(\langle f_{1}, \dots, f_{n} \rangle|y) & \text{Proportionality} \\ &= \underset{y}{\operatorname{argmax}} & \prod_{i=1}^{n} P_{F|Y}(f_{i}|y) & \text{Conditional independence} \\ &= \underset{y}{\operatorname{argmax}} & \sum_{i=1}^{n} \log \phi_{f_{i}}^{(y)} & \text{Monotonicity of logarithm} \end{split}$$

Instead of computing  $P_{Y|X}(y|x)$ 

- lacktriangle we compute  $P_{Y|F_1^n}(y|f_1^n)$  using some features  $f_1^n$  of x
- $\blacktriangleright$  instead of modelling  $P_{Y|F_1^n}$  directly, we modelled  $P_{F_1^n|Y}$  using tabular CPDs
- $\blacktriangleright$  and got to  $P_{Y|F_1^n}(y|f_1^n)$  via Bayes rule  $P_{Y|F_1^n}(y|f_1^n) \propto P_Y(y) P_{F_1^n|Y}(f_1^n|y)$

Instead of computing  $P_{Y|X}(y|x)$ 

- lacktriangle we compute  $P_{Y|F_1^n}(y|f_1^n)$  using some features  $f_1^n$  of x
- $\blacktriangleright$  instead of modelling  $P_{Y|F_1^n}$  directly, we modelled  $P_{F_1^n|Y}$  using tabular CPDs
- ▶ and got to  $P_{Y|F_1^n}(y|f_1^n)$  via Bayes rule  $P_{Y|F_1^n}(y|f_1^n) \propto P_Y(y)P_{F_1^n|Y}(f_1^n|y)$
- but isn't conditional independence a bit too strong?

Instead of computing  $P_{Y|X}(y|x)$ 

- lacktriangle we compute  $P_{Y|F_1^n}(y|f_1^n)$  using some features  $f_1^n$  of x
- $\blacktriangleright$  instead of modelling  $P_{Y|F_1^n}$  directly, we modelled  $P_{F_1^n|Y}$  using tabular CPDs
- ▶ and got to  $P_{Y|F_1^n}(y|f_1^n)$  via Bayes rule  $P_{Y|F_1^n}(y|f_1^n) \propto P_Y(y)P_{F_1^n|Y}(f_1^n|y)$
- but isn't conditional independence a bit too strong?

#### Consider the following example

Some parts are a bit slow and a little repetitive, but overall not too bad.

Instead of computing  $P_{Y|X}(y|x)$ 

- lacktriangle we compute  $P_{Y|F_1^n}(y|f_1^n)$  using some features  $f_1^n$  of x
- $\blacktriangleright$  instead of modelling  $P_{Y|F_1^n}$  directly, we modelled  $P_{F_1^n|Y}$  using tabular CPDs
- ▶ and got to  $P_{Y|F_1^n}(y|f_1^n)$  via Bayes rule  $P_{Y|F_1^n}(y|f_1^n) \propto P_Y(y)P_{F_1^n|Y}(f_1^n|y)$
- but isn't conditional independence a bit too strong?

Consider the following example

Some parts are a bit slow and a little repetitive, but overall not too bad.

It's riddled with generally negative words, but in the end the overall opinion is positive

Instead of computing  $P_{Y|X}(y|x)$ 

- lacktriangle we compute  $P_{Y|F_1^n}(y|f_1^n)$  using some features  $f_1^n$  of x
- $\blacktriangleright$  instead of modelling  $P_{Y|F_1^n}$  directly, we modelled  $P_{F_1^n|Y}$  using tabular CPDs
- ▶ and got to  $P_{Y|F_1^n}(y|f_1^n)$  via Bayes rule  $P_{Y|F_1^n}(y|f_1^n) \propto P_Y(y)P_{F_1^n|Y}(f_1^n|y)$
- but isn't conditional independence a bit too strong?

#### Consider the following example

Some parts are a bit slow and a little repetitive, but overall not too bad.

It's riddled with generally negative words, but in the end the overall opinion is positive

we need richer features, such as a bit slow, a little repetitive, not too bad

Instead of computing  $P_{Y|X}(y|x)$ 

- lacktriangle we compute  $P_{Y|F_1^n}(y|f_1^n)$  using some features  $f_1^n$  of x
- $\blacktriangleright$  instead of modelling  $P_{Y|F_1^n}$  directly, we modelled  $P_{F_1^n|Y}$  using tabular CPDs
- ▶ and got to  $P_{Y|F_1^n}(y|f_1^n)$  via Bayes rule  $P_{Y|F_1^n}(y|f_1^n) \propto P_Y(y)P_{F_1^n|Y}(f_1^n|y)$
- but isn't conditional independence a bit too strong?

#### Consider the following example

Some parts are a bit slow and a little repetitive, but overall not too bad.

It's riddled with generally negative words, but in the end the overall opinion is positive

- we need richer features, such as a bit slow, a little repetitive, not too bad
- **b** but an increase in feature space, e.g.  $O(v^3)$  for trigram features, leads to problems for parameter estimation

## Conditioning on high-dimensional data

#### The problem is that we only know tabular CPDs

If Y takes on values in  $\mathcal Y$  and X takes on values in  $\mathcal X$ , tabular CPDs associate a parameter  $\theta_y^{(x)}$  with each outcome y in context x

$$P_{Y|X}(y|x) = \operatorname{Cat}(y|\theta^{(x)}) = \theta_y^{(x)}$$

This can only work if  $|\mathcal{Y}|$  and  $|\mathcal{X}|$  are relatively small

▶ representation cost  $O(|\mathcal{Y}| \times |\mathcal{X}|)$ 

If x is itself very high-dimensional (e.g. a sentence), this cannot possibly work (as in this case  $\mathcal{X}\subseteq\Sigma^*$ )

Let's focus on the case where x is high-dimensional and y is binary How can we assign a value to  $P_{Y|X}(y|x)$ ?

- can we avoid a table lookup?
- can we let outcomes share statistical evidence?

Let's focus on the case where x is high-dimensional and y is binary How can we assign a value to  $P_{Y|X}(y|x)$ ?

- can we avoid a table lookup?
- can we let outcomes share statistical evidence?

Let's model the probability value!

Let's focus on the case where x is high-dimensional and y is binary How can we assign a value to  $P_{Y|X}(y|x)$ ?

- can we avoid a table lookup?
- can we let outcomes share statistical evidence?

#### Let's model the probability value!

lacktriangle suppose a function f(x,y) makes a finite summary of the aspects of the joint outcome (x,y) relevant to a problem

Let's focus on the case where x is high-dimensional and y is binary How can we assign a value to  $P_{Y|X}(y|x)$ ?

- can we avoid a table lookup?
- can we let outcomes share statistical evidence?

#### Let's model the probability value!

- lacktriangle suppose a function f(x,y) makes a finite summary of the aspects of the joint outcome (x,y) relevant to a problem
- we call  $f: \mathcal{X} \times \mathcal{Y} \to \mathbb{R}^D$  a feature function

Let's focus on the case where x is high-dimensional and y is binary How can we assign a value to  $P_{Y|X}(y|x)$ ?

- can we avoid a table lookup?
- can we let outcomes share statistical evidence?

#### Let's model the probability value!

- lacktriangle suppose a function f(x,y) makes a finite summary of the aspects of the joint outcome (x,y) relevant to a problem
- we call  $f: \mathcal{X} \times \mathcal{Y} \to \mathbb{R}^D$  a feature function
- $\blacktriangleright$  we can then make the probability value  $P_{Y|X}(y|x)$  depend functionally on f(x,y)

Let's focus on the case where x is high-dimensional and y is binary How can we assign a value to  $P_{Y|X}(y|x)$ ?

- can we avoid a table lookup?
- can we let outcomes share statistical evidence?

#### Let's model the probability value!

- lacktriangle suppose a function f(x,y) makes a finite summary of the aspects of the joint outcome (x,y) relevant to a problem
- we call  $f: \mathcal{X} \times \mathcal{Y} \to \mathbb{R}^D$  a feature function
- $\blacktriangleright$  we can then make the probability value  $P_{Y|X}(y|x)$  depend functionally on f(x,y)
- but we need to make sure that  $0 \leq P_{Y|X}(y|x) \leq 1$  and that  $\sum_y P_{Y|X}(y|x) = 1$

#### Feature function

An example of a binary feature function

Y	1	0
X	This film is fun and	This film is full of
	full of action	boring action
action <sub>+</sub>	1	0
action_	0	1
$boring_+$	0	0
boring_	0	1
full <sub>+</sub>	1	0
full_	0	1
fun <sub>+</sub>	1	0
fun_	0	0

Table: Feature function:  $f: \mathcal{X} \times \mathcal{Y} \to \{0,1\}^D$ 

▶ binary feature functions map the input to a *D*-dimensional vector of feature indicators

Suppose we have a D-dimensional real vector  $w \in \mathbb{R}^D$ 

Suppose we have a D-dimensional real vector  $w \in \mathbb{R}^D$ 

▶ Can we say  $P_{Y|X}(y|x) = w^{\top}f(x,y) = \sum_{d=1}^{D} w_d f_d(x,y)$  ?

Suppose we have a D-dimensional real vector  $w \in \mathbb{R}^D$ 

▶ Can we say  $P_{Y|X}(y|x) = w^{\top}f(x,y) = \sum_{d=1}^{D} w_d f_d(x,y)$  ? No! Because  $w^{\top}f(x,y)$  can be negative and  $w^{\top}f(x,y=1) + w^{\top}f(x,y=0)$  may not sum to 1

Suppose we have a D-dimensional real vector  $w \in \mathbb{R}^D$ 

- ▶ Can we say  $P_{Y|X}(y|x) = w^{\top}f(x,y) = \sum_{d=1}^{D} w_d f_d(x,y)$  ? No! Because  $w^{\top}f(x,y)$  can be negative and  $w^{\top}f(x,y=1) + w^{\top}f(x,y=0)$  may not sum to 1
- ▶ What if we make  $P_{Y|X}(y|x) = \exp\Big(w^{\top}f(x,y)\Big)$  ?

Suppose we have a D-dimensional real vector  $w \in \mathbb{R}^D$ 

- ► Can we say  $P_{Y|X}(y|x) = w^{\top}f(x,y) = \sum_{d=1}^{D} w_d f_d(x,y)$  ? No! Because  $w^{\top}f(x,y)$  can be negative and  $w^{\top}f(x,y=1) + w^{\top}f(x,y=0)$  may not sum to 1
- ▶ What if we make  $P_{Y|X}(y|x) = \exp\Big(w^{\top}f(x,y)\Big)$  ? Now we managed to ensure positivity, but still  $w^{\top}f(x,y=1) + w^{\top}f(x,y=0)$  may not sum to 1

Suppose we have a D-dimensional real vector  $w \in \mathbb{R}^D$ 

- ▶ Can we say  $P_{Y|X}(y|x) = w^{\top}f(x,y) = \sum_{d=1}^{D} w_d f_d(x,y)$  ? No! Because  $w^{\top}f(x,y)$  can be negative and  $w^{\top}f(x,y=1) + w^{\top}f(x,y=0)$  may not sum to 1
- ▶ What if we make  $P_{Y|X}(y|x) = \exp\left(w^{\top}f(x,y)\right)$  ? Now we managed to ensure positivity, but still  $w^{\top}f(x,y=1) + w^{\top}f(x,y=0)$  may not sum to 1
- ▶ The functional dependency on f(x,y) needs to be such that the result is a valid distribution, i.e. normalised across outcomes of Y|X=x

Suppose we have a D-dimensional real vector  $w \in \mathbb{R}^D$ 

- ▶ Can we say  $P_{Y|X}(y|x) = w^{\top}f(x,y) = \sum_{d=1}^{D} w_d f_d(x,y)$  ? No! Because  $w^{\top}f(x,y)$  can be negative and  $w^{\top}f(x,y=1) + w^{\top}f(x,y=0)$  may not sum to 1
- ▶ What if we make  $P_{Y|X}(y|x) = \exp\left(w^{\top}f(x,y)\right)$  ? Now we managed to ensure positivity, but still  $w^{\top}f(x,y=1) + w^{\top}f(x,y=0)$  may not sum to 1
- ▶ The functional dependency on f(x,y) needs to be such that the result is a valid distribution, i.e. normalised across outcomes of Y|X=x

$$P_{Y|X}(y|x) = \frac{\exp(w^{\top}f(x,y))}{\sum_{y' \in \mathcal{Y}} \exp(w^{\top}f(x,y'))}$$

We model the conditional using logistic regression

$$P_{Y|X}(y|x) = \frac{\exp(w^{\top}f(x,y))}{\sum_{y' \in \mathcal{Y}} \exp(w^{\top}f(x,y'))}$$

Then with a dataset  $\mathcal{D} = \{(x^{(k)}, y^{(k)})\}_{k=1}^N$  of i.i.d. observations, what's the maximum likelihood estimate for  $w \in \mathbb{R}^D$ ?

We model the conditional using logistic regression

$$P_{Y|X}(y|x) = \frac{\exp(w^{\top}f(x,y))}{\sum_{y' \in \mathcal{Y}} \exp(w^{\top}f(x,y'))}$$

Then with a dataset  $\mathcal{D} = \{(x^{(k)}, y^{(k)})\}_{k=1}^N$  of i.i.d. observations, what's the maximum likelihood estimate for  $w \in \mathbb{R}^D$ ?

Count and divide won't do ;)

We model the conditional using logistic regression

$$P_{Y|X}(y|x) = \frac{\exp(w^{\top}f(x,y))}{\sum_{y' \in \mathcal{Y}} \exp(w^{\top}f(x,y'))}$$

Then with a dataset  $\mathcal{D} = \{(x^{(k)}, y^{(k)})\}_{k=1}^N$  of i.i.d. observations, what's the maximum likelihood estimate for  $w \in \mathbb{R}^D$ ?

Count and divide won't do ;) recall where "count and divide" comes from

 $\blacktriangleright$  we looked for the solution to  ${\bf \nabla}_w {\cal L}(w|{\cal D})=0$ 

We model the conditional using logistic regression

$$P_{Y|X}(y|x) = \frac{\exp(w^{\top}f(x,y))}{\sum_{y' \in \mathcal{Y}} \exp(w^{\top}f(x,y'))}$$

Then with a dataset  $\mathcal{D} = \{(x^{(k)}, y^{(k)})\}_{k=1}^N$  of i.i.d. observations, what's the maximum likelihood estimate for  $w \in \mathbb{R}^D$ ?

We model the conditional using logistic regression

$$P_{Y|X}(y|x) = \frac{\exp(w^{\top}f(x,y))}{\sum_{y' \in \mathcal{Y}} \exp(w^{\top}f(x,y'))}$$

Then with a dataset  $\mathcal{D} = \{(x^{(k)}, y^{(k)})\}_{k=1}^N$  of i.i.d. observations, what's the maximum likelihood estimate for  $w \in \mathbb{R}^D$ ?

We look for w that is solution to  $\nabla_w \mathcal{L}(w|\mathcal{D}) = 0$  where

$$\mathcal{L}(w|\mathcal{D}) = \sum_{k=1}^{N} \underbrace{\log P_{Y|X}(y^{(k)}|x^{(k)}, w)}_{\ell(w|x^{(k)}, y^{(k)})}$$

is the log-likelihood function

# Let's start with a single training instance

The log-likelihood function gets a contribution  $\ell(w|x,y) = \log P_{Y|X}(y|x,w)$  from each training instance

Let's expand  $\ell$  slightly

$$\log P_{Y|X}(y|x, w) = \log \frac{\exp(w^{\top} f(x, y))}{\sum_{y' \in \mathcal{Y}} \exp(w^{\top} f(x, y'))}$$
=

# Let's start with a single training instance

The log-likelihood function gets a contribution  $\ell(w|x,y) = \log P_{Y|X}(y|x,w)$  from each training instance

Let's expand ℓ slightly

$$\log P_{Y|X}(y|x, w) = \log \frac{\exp(w^{\top} f(x, y))}{\sum_{y' \in \mathcal{Y}} \exp(w^{\top} f(x, y'))}$$
$$= w^{\top} f(x, y) - \log \sum_{\underline{y' \in \mathcal{Y}}} \exp(w^{\top} f(x, y'))$$
$$\underbrace{\sum_{Z(x|w)} \exp(w^{\top} f(x, y'))}_{Z(x|w)}$$

We need  $\nabla_w \log P_{Y|X}(y|x,w)$  but let's first take the gradient of the partition function  $\mathcal{Z}(x|w)$ 

Let 
$$\mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\Big(w^{\top} f(x,y)\Big)$$
,

Let 
$$\mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\Big(w^{\top} f(x,y)\Big)$$
, its gradient is 
$$\boldsymbol{\nabla}_{w} \mathcal{Z}(x|w) = \boldsymbol{\nabla}_{w} \sum_{y \in \mathcal{Y}} \exp\Big(w^{\top} f(x,y)\Big)$$

Let 
$$\mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\Bigl(w^{\top} f(x,y)\Bigr)$$
, its gradient is 
$$\nabla_w \mathcal{Z}(x|w) = \nabla_w \sum_{y \in \mathcal{Y}} \exp\Bigl(w^{\top} f(x,y)\Bigr)$$
$$= \sum_{y \in \mathcal{Y}} \nabla_w \exp\Bigl(w^{\top} f(x,y)\Bigr)$$

Let 
$$\mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\left(w^{\top} f(x,y)\right)$$
, its gradient is 
$$\nabla_w \mathcal{Z}(x|w) = \nabla_w \sum_{y \in \mathcal{Y}} \exp\left(w^{\top} f(x,y)\right)$$

$$= \sum_{y \in \mathcal{Y}} \nabla_w \exp\left(w^{\top} f(x,y)\right)$$

$$= \sum_{y \in \mathcal{Y}} \exp\left(w^{\top} f(x,y)\right) \nabla_w (w^{\top} f(x,y))$$

Let 
$$\mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\Bigl(w^{\top} f(x,y)\Bigr)$$
, its gradient is 
$$\nabla_w \mathcal{Z}(x|w) = \nabla_w \sum_{y \in \mathcal{Y}} \exp\Bigl(w^{\top} f(x,y)\Bigr)$$

$$= \sum_{y \in \mathcal{Y}} \nabla_w \exp\Bigl(w^{\top} f(x,y)\Bigr)$$

$$= \sum_{y \in \mathcal{Y}} \exp\Bigl(w^{\top} f(x,y)\Bigr) \nabla_w (w^{\top} f(x,y))$$

$$= \sum_{y \in \mathcal{Y}} \exp\Bigl(w^{\top} f(x,y)\Bigr) f(x,y)$$

Now we know that  $\mathbf{\nabla}_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\Bigl(w^{\top} f(x,y)\Bigr) f(x,y)$ 

Now we know that 
$$\nabla_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp(w^\top f(x,y)) f(x,y)$$

And 
$$\ell(w|x,y)=w^{\top}f(x,y)-\log\mathcal{Z}(x|w)$$
, thus its gradient is 
$$\nabla_w\ell(w|x,y)=$$

Now we know that 
$$\nabla_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp(w^\top f(x,y)) f(x,y)$$

And 
$$\ell(w|x,y) = w^{\top}f(x,y) - \log \mathcal{Z}(x|w)$$
, thus its gradient is 
$$\boldsymbol{\nabla}_w \ell(w|x,y) = \boldsymbol{\nabla}_w (w^{\top}f(x,y) - \log \mathcal{Z}(x|w))$$

Now we know that 
$$\nabla_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp(w^\top f(x,y)) f(x,y)$$

And 
$$\ell(w|x,y) = w^{\top} f(x,y) - \log \mathcal{Z}(x|w)$$
, thus its gradient is 
$$\nabla_w \ell(w|x,y) = \nabla_w (w^{\top} f(x,y) - \log \mathcal{Z}(x|w))$$
 
$$= f(x,y) - \nabla_w \log \mathcal{Z}(x|w)$$

Now we know that 
$$\nabla_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\Big(w^\top f(x,y)\Big) f(x,y)$$

And 
$$\ell(w|x,y) = w^{\top}f(x,y) - \log \mathcal{Z}(x|w)$$
, thus its gradient is 
$$\boldsymbol{\nabla}_{w}\ell(w|x,y) = \boldsymbol{\nabla}_{w}(w^{\top}f(x,y) - \log \mathcal{Z}(x|w))$$
$$= f(x,y) - \boldsymbol{\nabla}_{w}\log \mathcal{Z}(x|w)$$

Let's check the second term

$$\nabla_w \log \mathcal{Z}(x|w) =$$

Now we know that 
$$\nabla_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\Big(w^\top f(x,y)\Big) f(x,y)$$

And 
$$\ell(w|x,y) = w^{\top}f(x,y) - \log \mathcal{Z}(x|w)$$
, thus its gradient is 
$$\nabla_w \ell(w|x,y) = \nabla_w (w^{\top}f(x,y) - \log \mathcal{Z}(x|w))$$
$$= f(x,y) - \nabla_w \log \mathcal{Z}(x|w)$$

Let's check the second term

$$\nabla_w \log \mathcal{Z}(x|w) = \frac{1}{\mathcal{Z}(x|w)} \nabla_w \mathcal{Z}(x|w)$$

Now we know that 
$$\nabla_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\Big(w^\top f(x,y)\Big) f(x,y)$$

And 
$$\ell(w|x,y) = w^{\top}f(x,y) - \log \mathcal{Z}(x|w)$$
, thus its gradient is 
$$\nabla_w \ell(w|x,y) = \nabla_w (w^{\top}f(x,y) - \log \mathcal{Z}(x|w))$$
$$= f(x,y) - \nabla_w \log \mathcal{Z}(x|w)$$

Let's check the second term

$$\nabla_w \log \mathcal{Z}(x|w) = \frac{1}{\mathcal{Z}(x|w)} \nabla_w \mathcal{Z}(x|w)$$
$$= \frac{1}{\mathcal{Z}(x|w)} \sum_{y \in \mathcal{Y}} \exp(w^{\top} f(x,y)) f(x,y)$$

Now we know that 
$$\nabla_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\Big(w^\top f(x,y)\Big) f(x,y)$$

And 
$$\ell(w|x,y) = w^{\top}f(x,y) - \log \mathcal{Z}(x|w)$$
, thus its gradient is 
$$\nabla_w \ell(w|x,y) = \nabla_w (w^{\top}f(x,y) - \log \mathcal{Z}(x|w))$$
$$= f(x,y) - \nabla_w \log \mathcal{Z}(x|w)$$

Let's check the second term

$$\nabla_{w} \log \mathcal{Z}(x|w) = \frac{1}{\mathcal{Z}(x|w)} \nabla_{w} \mathcal{Z}(x|w)$$

$$= \frac{1}{\mathcal{Z}(x|w)} \sum_{y \in \mathcal{Y}} \exp(w^{\top} f(x, y)) f(x, y)$$

$$= \sum_{y \in \mathcal{Y}} \frac{\exp(w^{\top} f(x, y))}{\mathcal{Z}(x|w)} f(x, y)$$

Now we know that 
$$\nabla_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\Big(w^\top f(x,y)\Big) f(x,y)$$

And 
$$\ell(w|x,y) = w^{\top}f(x,y) - \log \mathcal{Z}(x|w)$$
, thus its gradient is 
$$\nabla_w \ell(w|x,y) = \nabla_w (w^{\top}f(x,y) - \log \mathcal{Z}(x|w))$$
 
$$= f(x,y) - \nabla_w \log \mathcal{Z}(x|w)$$

Let's check the second term

$$\nabla_{w} \log \mathcal{Z}(x|w) = \frac{1}{\mathcal{Z}(x|w)} \nabla_{w} \mathcal{Z}(x|w)$$

$$= \frac{1}{\mathcal{Z}(x|w)} \sum_{y \in \mathcal{Y}} \exp(w^{\top} f(x,y)) f(x,y)$$

$$= \sum_{y \in \mathcal{Y}} \frac{\exp(w^{\top} f(x,y))}{\mathcal{Z}(x|w)} f(x,y)$$

$$= \sum_{y \in \mathcal{Y}} P_{Y|X}(y|x,w) f(x,y)$$

Now we know that 
$$\nabla_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp\Big(w^\top f(x,y)\Big) f(x,y)$$

And 
$$\ell(w|x,y) = w^{\top}f(x,y) - \log \mathcal{Z}(x|w)$$
, thus its gradient is 
$$\boldsymbol{\nabla}_{w}\ell(w|x,y) = \boldsymbol{\nabla}_{w}(w^{\top}f(x,y) - \log \mathcal{Z}(x|w))$$
$$= f(x,y) - \boldsymbol{\nabla}_{w}\log \mathcal{Z}(x|w)$$

Let's check the second term

$$\nabla_{w} \log \mathcal{Z}(x|w) = \frac{1}{\mathcal{Z}(x|w)} \nabla_{w} \mathcal{Z}(x|w)$$

$$= \frac{1}{\mathcal{Z}(x|w)} \sum_{y \in \mathcal{Y}} \exp(w^{\top} f(x,y)) f(x,y)$$

$$= \sum_{y \in \mathcal{Y}} \frac{\exp(w^{\top} f(x,y))}{\mathcal{Z}(x|w)} f(x,y)$$

$$= \sum_{y \in \mathcal{Y}} P_{Y|X}(y|x,w) f(x,y) = \mathbb{E}[f(X,Y)|X = x;w]$$

## Putting everything together

#### We know

$$\mathbf{\nabla}_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp(w^{\top} f(x,y)) f(x,y)$$

$$\ell(w|x,y) = w^{\top} f(x,y) - \log \mathcal{Z}(x|w)$$

$$\nabla_w \ell(w|x,y) = f(x,y) - \nabla_w \log \mathcal{Z}(x|w)$$

$$\blacktriangleright$$
 and  $\nabla_w \log \mathcal{Z}(x|w) = \mathbb{E}[f(X,Y)|X=x;w]$ 

# Putting everything together

We know

$$\nabla_w \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp(w^{\top} f(x,y)) f(x,y)$$

$$\ell(w|x,y) = w^{\top} f(x,y) - \log \mathcal{Z}(x|w)$$

$$\nabla_w \ell(w|x,y) = f(x,y) - \nabla_w \log \mathcal{Z}(x|w)$$

▶ and 
$$\nabla_w \log \mathcal{Z}(x|w) = \mathbb{E}[f(X,Y)|X=x;w]$$

Then

$$\nabla_w \ell(w|x,y) = f(x,y) - \mathbb{E}[f(X,Y)|X=x;w]$$

Gradient descent

## Putting everything together

We know

$$\mathbf{\nabla}_{w} \mathcal{Z}(x|w) = \sum_{y \in \mathcal{Y}} \exp(w^{\top} f(x,y)) f(x,y)$$

$$\ell(w|x,y) = w^{\top} f(x,y) - \log \mathcal{Z}(x|w)$$

$$\nabla_w \ell(w|x,y) = f(x,y) - \nabla_w \log \mathcal{Z}(x|w)$$

▶ and 
$$\nabla_w \log \mathcal{Z}(x|w) = \mathbb{E}[f(X,Y)|X=x;w]$$

Then

$$\nabla_w \ell(w|x,y) = f(x,y) - \mathbb{E}[f(X,Y)|X=x;w]$$

There is no closed-form solution to  $\nabla_w \ell(w|x,y) = 0$ , but there is an iterative algorithm that converges to the solution

$$w^{(t+1)} = w^{(t)} + \gamma \nabla_{w^{(t)}} \ell(w^{(t)} | x, y)$$

 $\gamma > 0$  is called the *learning rate* (a hyperparameter)

Gradient descent

## Maximum likelihood estimation for logistic regression

We look for w that is solution to  $\nabla_w \mathcal{L}(w|\mathcal{D}) = 0$  where

$$\mathcal{L}(w|\mathcal{D}) = \sum_{k=1}^{N} \underbrace{\log P_{Y|X}(y^{(k)}|x^{(k)}, w)}_{\ell(w|x^{(k)}, y^{(k)})}$$

There is no closed-form solution  $\nabla_w \mathcal{L}(w|\mathcal{D})$ , but there is an iterative algorithm that converges to the solution

$$w^{(t+1)} = w^{(t)} + \gamma \underbrace{\sum_{k=1}^{N} \nabla_{w^{(t)}} \ell(w^{(t)} | x^{(k)}, y^{(k)})}_{\nabla_{w} \mathcal{L}(w | \mathcal{D})}$$

### Stochastic gradient ascent

We can use unbiased *stochastic gradient estimates* instead of the full gradient

$$w^{(t+1)} = w^{(t)} + \gamma^{(t)} \frac{M}{N} \sum_{s=1}^{M} \nabla_{w^{(t)}} \ell(w^{(t)} | x^{(s)}, y^{(s)})$$

where  $S \sim \mathcal{U}(1/N)$  selects training instances uniformly at random

## Stochastic gradient ascent

We can use unbiased *stochastic gradient estimates* instead of the full gradient

$$w^{(t+1)} = w^{(t)} + \gamma^{(t)} \frac{M}{N} \sum_{s=1}^{M} \nabla_{w^{(t)}} \ell(w^{(t)} | x^{(s)}, y^{(s)})$$

where  $S \sim \mathcal{U}(1/N)$  selects training instances uniformly at random The learning rate  $\gamma>0$  must follow a particular schedule, e.g.

$$\gamma^{(t)} = \frac{\gamma^{(t)}}{1 + \gamma^{(0)} \alpha t}$$

where the initial learning rate  $\gamma^{(0)}>0$  and the rate of decay  $\alpha>0$  are hyperparameters

## Regularisation

To avoid overfitting to training instances, we place a penalty on awkwardly large weights, our objective becomes

$$\underset{w \in \mathbb{R}^D}{\operatorname{argmax}} \ \mathcal{L}(w|\mathcal{D}) - \frac{\lambda}{2}||w||^2$$

where  $\lambda$  is the weight of the  $L_2$  regulariser

Our gradient becomes

$$\nabla_{w} \left( \mathcal{L}(w|\mathcal{D}) - \lambda ||w||^{2} \right) = \nabla_{w} \mathcal{L}(w|\mathcal{D}) - \frac{\lambda}{2} \nabla_{w} \sum_{d=1}^{D} w_{d}^{2}$$
$$= \nabla_{w} \mathcal{L}(w|\mathcal{D}) - \lambda \sum_{d=1}^{D} w_{d}$$

## Regularisation

To avoid overfitting to training instances, we place a penalty on awkwardly large weights, our objective becomes

$$\underset{w \in \mathbb{R}^D}{\operatorname{argmax}} \ \mathcal{L}(w|\mathcal{D}) - \frac{\lambda}{2}||w||^2$$

where  $\lambda$  is the weight of the  $L_2$  regulariser

Our gradient becomes

$$\nabla_{w} \left( \mathcal{L}(w|\mathcal{D}) - \lambda ||w||^{2} \right) = \nabla_{w} \mathcal{L}(w|\mathcal{D}) - \frac{\lambda}{2} \nabla_{w} \sum_{d=1}^{D} w_{d}^{2}$$
$$= \nabla_{w} \mathcal{L}(w|\mathcal{D}) - \lambda \sum_{d=1}^{D} w_{d}$$

Note that regularisation leads to dense gradients!

### Summary

Logistic regression allows us to express statistical dependencies between two variables through a finite set of features

- we can directly model a conditional probability using rich features of a high-dimensional conditioning context (this is called a *logistic cpd*)
- without the need for strong independence assumptions
- we have to estimate D parameters (the weights of a log-linear model)
- ► MLE does not have a closed-form solution, but gradient ascent gives us an iterative algorithm

Next class we will see how this can be used for various tasks e.g. sentiment classification, language identification, POS tagging, language modelling

### References I